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PROBABILITIES OF ONE-INCH SNOWFALL THRESHOLDS FOR THE UNITED STATES

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ABSTRACT

Following previous work, climatological statistics and computational methods are presented for obtaining the quantiles of first one-inch 24-hour snowfall. These are employed in computing the .05, .10, .30, .90 quantiles for 164 United States and Alaskan first-order Weather Bureau stations.

1. INTRODUCTION

The one-inch snowfall threshold is defined as the first day in fall or winter on which *one inch or more* of snowfall has occurred. This is expressed as a date; i. e., by day and month. It is a typical climatological threshold variable since it expresses the time at which some critical value is passed.

The climatological analysis for such snowfall thresholds was developed by Thom [1] and applied to various values of the threshold. He found that threshold date is approximately normally distributed for years in which a threshold value occurs. If the threshold value occurs every year or with a probability approaching one, the threshold date distribution is a normal distribution on day number counted from July 1. If the threshold does not occur every year, it was found that the threshold distribution could be expressed as a mixed distribution. This is a mixture of years with no threshold occurrence and years with threshold occurrence, the threshold occurrences being further distributed as to date. This theory led to expressions for the threshold quantiles and the mean threshold date.

The one-inch threshold quantiles are given for .05, .10, .30, and .90 probability for 164 cities in the United States and Alaska. These should be useful in many types of planning problems. In particular, they will be of interest

to snow tire manufacturers, snow removal equipment manufacturers, and those who are responsible for the removal of snow from streets and highways. The raw data employed in this study were originally compiled for a manufacturer of snow tires.

2. THE THRESHOLD DISTRIBUTION

The general threshold distribution function was given in [1] as

$$G(t) = pF(t) + q(t > 365) \quad (1)$$

Here $G(t)$ is the probability of a threshold occurrence before date t , p is the probability of a year with a threshold occurrence, $F(t)$ is the probability of a threshold occurrence before t in years which have threshold occurrences, and $q(t > 365)$ is the probability of no threshold occurrence and is included only after $F(t)$ approaches one. The non-occurrence of a threshold here is formally equivalent to the occurrence of a threshold outside the dated threshold population, or outside the 365-day year allowed for thresholds to occur and therefore not considered as an occurrence.

Although the result is the same, the threshold distribution may be developed in a somewhat different fashion as follows: According to the development above, $F(t)$ is the probability of a threshold occurring before t , hence, $1 - F(t)$ is the probability of a threshold after t . We may

now express the distribution function for a threshold after t as

$$H(t) = p[1 - F(t)] + q. \quad (2)$$

Hence, the distribution function for a threshold before t is

$$G(t) = 1 - H(t) = 1 - p[1 - F(t)] - q = pF(t) \quad (3)$$

This becomes identical with equation (1) when we add q , the probability of no threshold occurring, and meets the condition that a distribution function must approach unity as t increases without limit. Equation (3) may also be looked upon as a conditional probability; i. e., the product of the probability of a threshold occurring by the probability that it will occur before t , given that a threshold has occurred. Graphical illustrations of the frequency distribution and distribution function are given in [1]. Equation (3) will be found convenient for the computation of the quantiles.

In reference [1] the hypothesis that $F(t)$ is a normal distribution function was examined and the hypothesis of normality found to be well met by the samples studied. Similar procedures were applied in examining the normality hypothesis for this much larger sample of stations on the one-inch threshold. As was to be expected, statistical analysis showed no general significance departure from normality. Hence, we have again assumed t to be approximately normally distributed with estimated mean \bar{t} and standard deviation s .

3. COMPUTATION OF THE QUANTILES

Since the quantiles are inverse functions of the variate, they may be obtained by inverting equation (3). Writing $N(t)$ for $F(t)$ to indicate a normal distribution, equation (3) becomes

$$N(t) = G(t)/p.$$

Inverting the function on the left gives

$$t = N^{-1}[G(t)/p] \quad (4)$$

In order to use the standard normal tables, t must be converted to the standardized variate $x = (t - \bar{t})/s$. Substituting this in (4) we find

$$(t - \bar{t})/s = N^{-1}[G(x)/p].$$

Hence

$$t_G = sN^{-1}[G(x)/p] + \bar{t} \quad (5)$$

where t_G is a quantile with probability $G(x)$.

Using the climatological series of one-inch thresholds the mean \bar{t} and the standard deviation s were estimated

in the usual fashion. Since p is the probability of such a threshold occurrence, it was estimated by the ratio of the number of years with thresholds to the total number of years of record n . These are shown for each station in table 1.

The procedure for obtaining the quantiles may be best illustrated by an example. Taking the first station in the table, Anniston, Ala., we find $\bar{t} = 2/2$ (i. e., Feb. 2), $s = 32.1$, and $p = 0.2889$. Converting \bar{t} to day number gives 217 (the exact value was 216.6). For the .05 quantile $G(x) = .05$. Substituting these values in equation (5) yields

$$\begin{aligned} t_{.05} &= 32.1N^{-1}(.05/.2889) + 216.6 \\ &= 32.1N^{-1}(.173) + 216.6 \end{aligned}$$

Referring to a standard normal table and keeping in mind that .173 is a probability, we find $N^{-1}(.173) = -0.94$. Substituting this value yields

$$t_{.05} = 32.1(-.94) + 216.6 = 186.$$

Converting this day number to date we find $t_{.05} = 1/2$, or January 2, which agrees with table 1. It is clear that any quantile may be readily obtained by this procedure.

It should be noted that since p is the amount of probability in the actual occurrence of a threshold, the quantile probabilities cannot exceed this amount. Hence, when the computed quantile probability exceeds p there is no entry in the table. No stations were included where the probability of a threshold occurrence was less than 0.10. Each quantile $t_{.05}$, etc., of table 1 is the date for which the probability of one inch or more in 24 hours occurring earlier in date is the quantile probability. Thus, for Burlington, Iowa, as seen in the table, the probability of a threshold occurrence before 10/23 is 0.05, before 11/2 is 0.10, before 11/21 is 0.30 and before 1/6 is 0.90. Also in other terms, the first one-inch 24-hour snowfall will occur on the average once in 20 years before 10/23, once in 10 years before 11/2, about once in 3 years before 11/21, and in nine out of ten years before 1/6.

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REFERENCE

1. H. C. S. Thom, "Climatological Analysis of Snowfall Thresholds," *Archiv für Meteorologie, Geophysik, und Bioklimatologie, Serie B*, vol. 8, No. 2, 1957, pp. 195-201.

TABLE 1.—Statistics on one-inch threshold for snowfall for 164 first-order stations, listed alphabetically by states. The entries under \bar{t} and t_q indicate "month/date"

Station	\bar{t}	s	n	p	$t_{.05}$	$t_{.10}$	$t_{.25}$	$t_{.50}$	Station	\bar{t}	s	n	p	$t_{.05}$	$t_{.10}$	$t_{.25}$	$t_{.50}$
Anniston, Ala.	2/2	32.1	45	0.2889	1/2	1/20			Binghamton, N. Y.	11/26	15.8	51	1.0000	10/31	11/5	11/17	12/16
Birmingham, Ala.	1/24	35.3	51	.4118	12/14	12/31	2/15		Buffalo, N. Y.	11/19	17.6	51	1.0000	10/21	10/28	11/10	12/12
Tucson, Ariz.	1/10	25.0	21	.3810	12/13	12/25	1/30		Canton, N. Y.	11/13	18.6	44	1.0000	10/13	10/20	11/3	12/7
Fort Smith, Ark.	1/19	29.2	51	.7843	12/6	12/17	1/10		New York, N. Y.	12/19	17.5	51	.9804	11/20	11/27	12/10	1/12
Little Rock, Ark.	1/15	33.2	51	.6471	11/29	12/12	1/12		Oswego, N. Y.	11/22	18.9	51	1.0000	10/30	11/4	11/15	12/10
Eureka, Calif.	1/23	25.9	51	.1176	1/18	2/19			Rochester, N. Y.	11/21	14.5	51	1.0000	10/28	11/3	11/14	12/10
Red Bluff, Calif.	1/6	18.8	51	.4118	12/15	12/24	1/18		Syracuse, N. Y.	11/18	18.6	49	1.0000	10/24	10/29	11/9	12/3
Denver, Colo.	10/27	18.5	51	1.0000	9/26	10/3	10/17	11/19	Asheville, N. C.	12/26	30.7	49	.8980	11/8	11/19	12/13	
Grand Junction, Colo.	12/2	23.3	51	.9804	10/25	11/2	11/20	1/3	Charlotte, N. C.	1/16	32.6	51	.6863	11/30	12/13	1/11	
Pueblo, Colo.	11/16	24.2	51	1.0000	10/8	10/16	11/4	12/17	Greensboro, N. C.	1/7	29.1	23	.6522	11/27	12/9	1/4	
Hartford, Conn.	12/9	19.4	47	1.0000	11/7	11/14	11/29	1/3	Hatteras, N. C.	1/12	24.5	51	.2157	12/25	1/9		
New Haven, Conn.	12/11	20.5	51	1.0000	11/7	11/15	11/30	1/6	Raleigh, N. C.	1/12	30.6	51	.7451	11/28	12/9	1/5	
Wilmington, Del.	12/22	25.6	43	.9535	11/11	11/20	12/10	2/1	Wilmington, N. C.	1/20	29.5	51	.2941	12/23	1/8		
Atlanta, Ga.	1/17	34.8	51	.2745	12/17	1/5			Bismarck, N. Dak.	11/7	26.2	51	1.0000	9/25	10/5	10/25	12/11
Augusta, Ga.	1/17	36.9	51	.1176	1/10	2/24			Devils Lake, N. Dak.	11/9	20.4	46	1.0000	10/7	10/14	10/30	12/5
Macon, Ga.	1/22	37.7	51	.1569	1/5	2/5			Fargo, N. Dak.	11/17	25.1	51	1.0000	10/7	10/16	11/4	12/19
Boise, Idaho.	12/6	21.3	51	.9804	11/1	11/9	11/25	1/4	Williston, N. Dak.	11/11	26.7	51	1.0000	9/28	10/8	10/28	12/15
Pocatello, Idaho.	11/21	25.0	51	1.0000	10/11	10/20	11/8	12/23	Akron-Canton, Ohio.	11/27	21.1	20	1.0000	10/23	10/31	11/16	12/24
Cañero, Ill.	12/30	29.5	51	.9412	11/12	11/23	12/16	2/19	Cincinnati, Ohio.	11/26	23.5	51	.9804	11/8	11/17	12/4	1/18
Chicago, Ill.	12/6	18.7	51	1.0000	11/5	11/12	11/26	12/30	Cleveland, Ohio.	12/18	19.2	51	1.0000	10/27	11/3	11/18	12/22
Moline, Ill.	12/11	21.0	20	.9500	11/7	11/14	12/1	1/14	Columbus, Ohio.	12/13	25.6	51	.9804	11/1	11/11	11/30	1/18
Peoria, Ill.	12/12	24.5	45	1.0000	11/1	11/10	11/29	1/12	Dayton, Ohio.	12/13	26.1	40	.9750	11/1	11/10	11/30	1/19
Springfield, Ill.	12/17	27.8	51	1.0000	11/1	11/11	12/2	1/21	Sandusky, Ohio.	12/8	20.7	51	.9804	11/4	11/12	11/27	1/6
Evansville, Ind.	12/24	28.5	51	.9020	11/9	11/20	12/12	3/16	Toledo, Ohio.	12/7	18.5	51	1.0000	11/7	11/14	11/28	12/31
Fort Wayne, Ind.	12/7	23.0	40	.9750	10/31	11/8	11/26	1/9	Oklahoma City, Okla.	1/8	31.3	51	.8627	11/20	12/1	12/26	
Indianapolis, Ind.	12/15	23.3	51	1.0000	11/7	11/15	12/3	1/14	Tulsa, Okla.	12/23	15.6	12	1.0000	10/2	12/8	12/20	1/17
Terre Haute, Ind.	12/19	24.1	30	1.0000	11/9	11/18	12/6	1/19	Baker, Oreg.	11/23	19.0	50	1.0000	12/23	10/30	11/13	12/17
Burlington, Iowa.	12/4	25.7	49	1.0000	10/23	11/2	11/21	1/6	Portland, Oreg.	1/10	29.8	51	.7843	11/26	12/7	1/1	
Charles City, Iowa.	11/26	20.1	51	1.0000	10/24	10/31	11/15	12/22	Medford, Oreg.	1/4	26.5	40	.6750	11/26	12/7	12/31	
Davenport, Iowa.	12/8	23.8	51	1.0000	10/30	11/8	11/26	1/8	Roseburg, Oreg.	1/11	21.8	51	.6078	12/12	12/21	1/11	
Des Moines, Iowa.	12/5	21.9	51	1.0000	10/30	11/7	11/24	1/2	Erie, Pa.	11/13	16.2	51	1.0000	10/17	10/23	11/5	12/4
Dubuque, Iowa.	12/1	20.2	51	1.0000	10/28	11/5	11/20	12/26	Harrisburg, Pa.	12/13	20.6	51	1.0000	11/9	11/16	12/2	1/8
Keokuk, Iowa.	12/13	22.2	47	1.0000	11/6	11/14	12/1	1/10	Philadelphia, Pa.	12/22	24.1	51	1.0000	11/13	11/22	12/10	1/22
Sioux City, Iowa.	11/24	21.1	51	1.0000	10/21	10/28	11/13	12/21	Pittsburgh, Pa.	12/3	22.2	51	1.0000	10/28	11/5	11/22	1/1
Concordia, Kans.	12/15	32.8	51	.9412	10/23	11/4	11/30	2/9	Reading, Pa.	12/13	18.4	47	1.0000	11/13	11/19	12/3	1/6
Dodge City, Kans.	12/10	28.9	51	.9608	10/24	11/4	11/28	1/23	Seranton, Pa.	12/3	18.8	51	1.0000	11/2	11/9	11/23	12/27
Topeka, Kans.	12/12	27.5	51	1.0000	10/28	11/7	11/28	1/17	Block Island, R. I.	12/28	24.0	51	1.0000	11/18	11/27	12/15	1/27
Wichita, Kans.	12/18	24.4	51	.9608	11/9	11/17	12/6	1/25	Providence, R. I.	12/15	21.1	47	1.0000	11/10	11/18	12/4	1/11
Lexington, Ky.	12/15	25.4	51	.9804	11/3	11/12	12/2	1/19	Columbia, S. C.	1/11	36.6	51	.2157	12/15	1/7		
Louisville, Ky.	12/22	28.1	51	.9412	11/6	11/17	12/9	2/8	Greenville, S. C.	1/16	25.8	34	.5588	12/12	12/23	1/18	
Shreveport, La.	1/15	25.8	51	.3333	12/19	1/2	2/17	12/17	Huron, S. Dak.	11/22	26.0	51	.9804	10/11	10/20	11/9	12/28
Eastport, Maine.	11/25	16.8	51	1.0000	10/29	11/4	11/17	12/19	Rapid City, S. Dak.	11/8	25.7	51	1.0000	9/27	10/6	10/26	12/11
Portland, Maine.	11/30	14.7	51	1.0000	11/6	11/11	11/22	12/19	Chatanooga, Tenn.	1/5	28.6	51	.6471	11/26	12/7	1/3	
Baltimore, Md.	12/26	30.6	51	.9804	11/6	11/17	12/10	2/7	Knoxville, Tenn.	1/3	33.1	51	.8627	11/12	11/25	12/21	
Washington, D. C.	12/22	28.9	51	1.0000	11/5	11/15	12/7	1/28	Nashville, Tenn.	1/6	34.8	51	.8431	11/13	11/26	12/25	
Boston, Mass.	12/14	18.1	51	1.0000	11/14	11/21	12/5	1/6	Akron, Tex.	1/15	31.6	51	.6471	12/1	12/14	1/12	
Nantucket, Mass.	12/25	21.6	51	1.0000	11/19	11/27	12/14	1/21	Amarillo, Tex.	12/10	33.5	51	.9412	11/17	10/29	11/24	2/5
Alpena, Mich.	11/17	16.4	51	1.0000	10/21	10/27	11/8	12/8	Anstis, Tex.	1/22	33.6	50	.1800	1/3	1/18		
Detroit, Mich.	12/1	17.1	51	1.0000	11/3	11/9	11/22	12/23	Dallas, Tex.	1/11	27.4	38	.5000	12/7	12/19	1/18	
Escanaba, Mich.	11/21	17.3	51	1.0000	10/24	10/30	11/12	12/13	Del Rio, Tex.	1/17	23.0	45	.2000	1/1	1/17		
Grand Rapids, Mich.	11/21	14.3	48	1.0000	10/28	11/2	11/13	12/9	El Paso, Tex.	1/4	27.3	51	.5490	11/29	12/11	1/8	
Lansing, Mich.	11/22	17.0	41	1.0000	10/25	10/31	11/13	12/14	Fort Worth, Tex.	1/13	26.0	51	.5490	12/9	12/20	1/16	
Marquette, Mich.	11/2	15.3	51	1.0000	10/8	10/14	10/25	11/22	Salt Lake City, Utah.	11/15	18.7	51	1.0000	10/16	10/23	11/6	12/9
Sault Ste Marie, Mich.	11/3	13.3	51	1.0000	10/12	10/17	10/27	11/20	Modena, Utah.	11/26	25.8	40	1.0000	10/15	10/24	11/13	12/29
Duluth, Minn.	11/13	21.2	51	1.0000	10/9	10/17	11/2	12/10	Burlington, Vt.	11/18	14.9	55	1.0000	10/25	10/30	11/11	12/7
Minneapolis, Minn.	11/20	22.3	51	1.0000	10/15	10/23	11/9	12/19	Northfield, Vt.	11/11	17.1	43	1.0000	10/14	10/21	11/3	12/3
Meridian, Miss.	1/29	30.2	51	.2353	1/5	1/23			Cape Henry, Va.	1/16	25.2	51	.7647	12/9	12/19	1/9	
Vicksburg, Miss.	1/20	23.1	51	.3333	12/27	1/8	2/19		Lynchburg, Va.	1/2	30.8	51	.9804	11/12	11/24	12/17	2/13
Columbia, Mo.	12/19	29.3	51	.9804	10/31	11/11	12/4	1/28	Norfolk, Va.	1/12	27.2	51	.8431	12/1	12/11	1/2	
Kansas City, Mo.	12/12	27.7	51	1.0000	10/27	11/6	11/27	1/16	Richmond, Va.	1/2	28.7	51	.9216	11/17	11/28	12/20	2/28
St. Joseph, Mo.	12/16	32.5	41	.9756	10/24	11/5	11/30	1/31	North Head, Wash.	1/17	28.2	49	.5714	12/10	12/22	1/19	
St. Louis, Mo.	12/21	23.6	51	.9804	11/12	11/21	12/9	1/22	Seattle, Wash.	1/2	29.3	51	.7255	11/20	12/2	12/27	
Springfield, Mo.	12/22	29.4	51	1.0000	11/4	11/14	12/7	1/29	Spokane, Wash.	11/27	22.1	51	1.0000	10/21	10/29	11/15	12/25
Billings, Mont.	11/3	24.1	38	1.0000	9/24	10/3	10/21	12/3	Tacoma, Wash.	12/26	30.1	51	.8039	11/10	11/22	12/17	
Butte, Mont.	10/10	23.4	12	1.0000	9/2	9/10	10/6	11/9	Tatoosie Island, Wash.	1/2	28.2	48	.6250	11/23	12/5	1/1	
Great Falls, Mont.	10/19	25.0	11	1.0000	9/8	9/17	10/13	12/7	Walla Walla, Wash.	12/15	26.8	51	.9608	11/1	11/11	12/2	1/25
Havre, Mont.	10/29	30.5	51	1.0000	9/9	9/20	10/6	11/20	Yakima, Wash.	12/11	23.7	32	1.0000	11/2	11/11	11/29	1/10
Helena, Mont.	10/19	24.8	51	1.0000	9/8	9/17	10/6	11/20	Elkins, W. Va.	11/23	24.4	51	1.0000	10/14	10/23	11/10	12/24
Kalispell, Mont.	11/14	23.1	51	1.0000	10/7	10/15	11/2	12/13	Parkersburg, W. Va.	12/12	25.4	51	1.0000	11/1	11/10	11/29	1/14
Missoula, Mont.	11/24	26.6	51	.9804	10/11	10/21	11/10	12/31	Green Bay, Wis.	11/27	21.4	51	1.0000	10/23	10/30	11/16	12/24
Lincoln, Nebr.	12/7	25.5	51	1.0000	10/26	11/4	11/10	12/29	La Crosse, Wis.	11/26	19.7	51	1.0000	10/24	10/31	11/15	12/21
North Platte, Nebr.	11/24	27.5	51	1.0000													